

# Discussion section 9

Math 181B

## 1 Review

### 1.1 The Sign test

The Sign test is a non-parametric test whose null hypothesis is about checking if the median of a distribution is equal to some specific value. The idea is simple: The number of samples greater than the median should have a binomial distribution.

More specifically, let  $y_1, y_2, \dots, y_n$  be a random sample of size  $n$  from any continuous distribution having median  $\tilde{\mu}$ , where  $n \geq 10$ . Let  $k$  denote the number of  $y_i$  's greater than  $\tilde{\mu}_0$ , and let  $z = \frac{k-n/2}{\sqrt{n/4}}$ .

- To test  $H_0 : \tilde{\mu} = \tilde{\mu}_0$  versus  $H_1 : \tilde{\mu} > \tilde{\mu}_0$  at the  $\alpha$  level of significance, reject  $H_0$  if  $z \geq z_\alpha$ .
- To test  $H_0 : \tilde{\mu} = \tilde{\mu}_0$  versus  $H_1 : \tilde{\mu} < \tilde{\mu}_0$  at the  $\alpha$  level of significance, reject  $H_0$  if  $z \leq -z_\alpha$ .
- To test  $H_0 : \tilde{\mu} = \tilde{\mu}_0$  versus  $H_1 : \tilde{\mu} \neq \tilde{\mu}_0$  at the  $\alpha$  level of significance, reject  $H_0$  if  $z$  is either (1)  $\leq -z_{\alpha/2}$  or (2)  $\geq z_{\alpha/2}$ .

**When**  $n < 10$ , the normal approximation to binomial is not appropriate. Therefore, we need to use the exact binomial distribution to compute the  $p$ -value and then make a conclusion about the hypothesis test.

### 1.2 Wilcoxon Tests

#### 1.2.1 Wilcoxon Signed Rank Test

Wilcoxon test is another non-parametric test that can be used to test location (mean of one sample, mean of two samples), etc. If we define the Wilcoxon signed rank statistic,  $w$  as

$$w = \sum_{i=1}^n r_i z_i;$$

where

$$z_i = \begin{cases} 0 & \text{if } y_i - \mu_0 < 0 \\ 1 & \text{if } y_i - \mu_0 > 0 \end{cases},$$

and  $r_i$  are ranks of  $|y_1 - \mu_0|, |y_2 - \mu_0|, \dots, |y_n - \mu_0|$ , where  $\mu_0$  is the mean under the null hypothesis.

If  $y_1, y_2, \dots, y_n$  be a set of independent observations drawn, respectively, from the continuous and symmetric (but not necessarily identical) pdfs  $f_{Y_i}(y), i = 1, 2, \dots, n$ . Suppose that each of the  $f_{Y_i}(y)$  's has the same mean  $\mu$ . Under the null hypothesis  $H_0 : \mu = \mu_0$ ,  $p_W(w)$  can be calculated as

$$p_W(w) = P(W = w) = \left( \frac{1}{2^n} \right) \cdot c(w)$$

where  $c(w)$  is the coefficient of  $e^{wt}$  in the expansion of

$$\prod_{i=1}^n (1 + e^{it}).$$

**When the sample size is large**  $n > 12$ , we can get a normal approximation, i.e.,

$$\frac{W - [n(n+1)]/4}{\sqrt{[n(n+1)(2n+1)]/24}}$$

can be adequately approximated by a normal distribution.

Therefore, we can accomplish the hypothesis test based on the exact probability mass function  $p_W(w)$  or the normal distribution. For example, when the sample size is large, let

$$z = \frac{w - [n(n+1)]/4}{\sqrt{[n(n+1)(2n+1)]/24}}$$

- To test  $H_0 : \mu = \mu_0$  versus  $H_1 : \mu > \mu_0$  at the  $\alpha$  level of significance, reject  $H_0$  if  $z \geq z_\alpha$ .
- To test  $H_0 : \mu = \mu_0$  versus  $H_1 : \mu < \mu_0$  at the  $\alpha$  level of significance, reject  $H_0$  if  $z \leq -z_\alpha$ .
- To test  $H_0 : \mu = \mu_0$  versus  $H_1 : \mu \neq \mu_0$  at the  $\alpha$  level of significance, reject  $H_0$  if  $z$  is either (1)  $\leq -z_{\alpha/2}$  or (2)  $\geq z_{\alpha/2}$ .

### 1.2.2 Wilcoxon Rank Sum Test

Wilcoxon Rank Sum Test can be used to test  $H_0 : \mu_x = \mu_y$  for two continuous distributions which have the same shape and the same standard deviation, but they may differ with respect to location.

More specifically, let  $x_1, x_2, \dots, x_n$  and  $y_{n+1}, y_{n+2}, \dots, y_{n+m}$  be two independent random samples from  $f_X(x)$  and  $f_Y(y)$ , respectively, where the two pdfs are the same except for a possible shift in location. Let  $r_i$  denote the rank of the  $i$ th observation in the combined sample (where the smallest observation is assigned a rank of 1 and the largest observation, a rank of  $n+m$ ). Let

$$w' = \sum_{i=1}^{n+m} r_i z_i$$

where  $z_i$  is 1 if the  $i$ -th observation comes from  $f_X(x)$  and 0, otherwise. Then

$$E(W') = \frac{n(n+m+1)}{2}$$

$$\text{Var}(W') = \frac{nm(n+m+1)}{12}$$

and

$$\frac{W' - n(n+m+1)/2}{\sqrt{nm(n+m+1)/12}}$$

is approximately a standard normal distribution if  $n > 10$  and  $m > 10$ .